



The demand for punishment

Jeffrey P. Carpenter*

Department of Economics, Munroe Hall, Middlebury College, Middlebury, VT 05753, USA

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Abstract

While many experiments demonstrate that behavior differs from the predictions of traditional economic theory, they have not shown that economic reasoning is necessarily incorrect. Instead, these experiments illustrate that the preferences of homo economicus have been mis-specified. Modeled with social preferences, it may be rational for agents to forego material gains. Social dilemmas are examples in which punishment is not credible and yet people often pay to reprimand other participants. At the same time, we show that these people also react to changes in the price of punishing and income as if punishment was an ordinary and inferior good.

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1. Introduction

At this point in the evolution of experimental and behavioral economics, laboratory experiments have provided more new questions about economic behavior than answers. Instead of confirming the standard tenets of neoclassical economics, experiments have identified decision-making anomalies (Camerer, 1995), preference reversals (Tversky et al., 1990), and non-standard or “social” preferences (Camerer and Fehr, 2001; Carpenter, 2002; Charness and Rabin, 2002).¹ Expanding on the idea of social preferences, experiments have shown that instead of being selfish and myopic, average participants are much better described as trusting and trustworthy (Berg et al., 1995), fair (Güth et al., 1982; Fehr et al., 1993), and cooperative (Isaac et al., 1984), but, as

* Tel.: +1 802 443 3241; fax: +1 802 443 2084.

E-mail address: jpc@middlebury.edu (J.P. Carpenter).

¹ In fact, competitive markets is one of the few areas where experiments have come close to confirming existing theories (Davis and Holt, 1993).

28 motivation for what follows, it is important to stress that they can also be vindictive (Camerer and
29 Thaler, 1995; Fehr and Gächter, 2000).

30 However, the fact that many economic models predict behavior that is at odds with what we
31 observe in experiments may be because we have misspecified peoples' preferences, not because
32 the methodology of economics is fundamentally flawed. Although people behave as if they have
33 preferences for cooperation and retaliation, they may still react to incentives in ways predicted
34 by standard economic logic. For example, if we hypothesize that peoples' observed preferences
35 for cooperation operate like preferences for more standard consumption goods, then we might
36 expect people to cooperate less when the implied price of cooperation increases, just as they tend
37 to buy fewer ordinary goods when the price increases.

38 In addition to being predisposed to cooperate, recent experiments have demonstrated that
39 people retaliate against perceived injustices, even when doing so is costly and the material benefits
40 of doing so are small or nonexistent. This evidence (reviewed below) leads one to believe that
41 many participants have a preference for punishing asocial behavior. In the experiment reported on
42 herein, we test, in a controlled setting, whether such a nonstandard preference behaves according
43 to standard economic reasoning.

44 This research is unique because it is among the first to examine explicitly whether standard
45 economic tools can explain behavior motivated by the nonstandard preference to punish free
46 riders. At the same time, this research is linked to other recent work in behavioral economics. One
47 area of research examines the sacrifices that people are willing to endure to assure fair outcomes
48 and, in this sense, examines the price responsiveness of fairness preferences. Examples of this
49 literature include Eckel and Grossman (1996), Suleiman (1996), and Zwick and Chen (1999). In a
50 second related project, Andreoni and co-workers (Andreoni and Vesterlund, 2001; Andreoni and
51 Miller, 2002; Andreoni et al., 2003) empirically recover utility functions that are based on social
52 preferences. Variants of these utility functions could, in principle, generate the sort of demand for
53 punishment functions that we estimate below.

54 We proceed by briefly reviewing the literature on social dilemma experiments in which players
55 were given the opportunity to punish each other. Hopefully, this review will convince the reader
56 that cooperation and retaliation are robust behaviors. We then discuss the current experiment
57 that was designed to examine whether peoples' preferences for punishment behave according
58 to standard economic logic. Specifically, the experiment provides us with data which we use to
59 estimate the demand for punishment. Our analysis indicates that punishment is both ordinary and
60 inferior, but is also relatively inelastic with respect to both price and income.

61 **2. Fairness, cooperation, and punishment**

62 The first evidence of a preference for punishing asocial behavior came from one-shot ultimatum
63 games in which a first-mover makes an offer to share a sum of money with a second-mover who
64 accepts or rejects this offer (Güth et al., 1982; Camerer, 1995). Although any division of the pie
65 can be supported as an equilibrium of this game, subgame perfection leads one to expect that the
66 first-mover will receive all (or almost all) of the money because selfish second-movers will always
67 accept small offers rather than reject them and get nothing. Despite this unambiguous prediction,
68 nearly all small offers are rejected, and the most common explanation given by second-movers is
69 that they are retaliating against greedy first-movers (Pillutla and Murnighan, 1996).

70 Punitive behavior has also been witnessed in social dilemma games in which individual and
71 group incentives are at odds, and therefore, free-riding is expected from selfish players. One of the
72 first of these experiments was conducted by Ostrom et al. (1992). In this common pool resource

73 experiment players cooperate with each other by not extracting too much from an open-access
74 and subtractable resource. Resource use is problematic because by extracting, one player imposes
75 a negative externality on all the other players. Under these incentives, the authors showed that
76 when costly punishment was allowed, cooperative players used it to regulate the behavior of
77 over-extractors (i.e., free riders) and the gross efficiency of extraction increased, especially when
78 communication was allowed too.

79 Considering positive rather than negative externalities, [Fehr and Gächter \(2000\)](#) tested whether
80 costly punishment could curtail free riding in a public goods experiment. In the voluntary con-
81 tribution mechanism players emit a positive externality every time they contribute to a group
82 project, the benefits of which are shared by the entire group. Given this structure, selfish players
83 should contribute nothing and free ride on the contributions of others. Fehr and Gächter's results
84 mirror those of [Ostrom et al.](#) in that they find that many contributors are willing to pay to punish
85 those who contribute less than the average. Further, the (theoretically incredible) threat to punish
86 reduces free riding dramatically. These results suggest that when subjects punish free riders they
87 are expressing a social preference for retaliation because they punish despite having to pay to do
88 so and despite the negligible material benefits that are expected to follow punishment.

89 There are many recent extensions of the Fehr and Gächter results. For example, [Bochet et al. \(in](#)
90 [press\)](#) confirm that punishment increases contributions, but they also find that it does not increase
91 them as much or as efficiently (when one accounts for the payoff consequences to both punisher
92 and target) as communication. [Bowles et al. \(2001\)](#) also find increases in contributions due to
93 punishment and extend the literature by looking at the implications of group size and the return
94 on the public good. Their data suggest that the amount of punishment received by free riders is
95 increasing in both the size of the group and the return on the public good. The end result is that
96 large groups that generate large externalities contribute at very high levels when punishment is
97 allowed.

98 In an extension of [Bowles et al. \(2001\)](#), [Carpenter \(in press\)](#) examines the interaction of
99 group size, public good productivity, and monitoring technology and finds that punishment (as
100 a contribution elicitation mechanism) is sensitive to the structure of groups. When each group
101 member can monitor and punish all the other members of her group, contributions are high
102 regardless of group size and the return on the public good. However, the most efficient use of
103 punishment occurs when group members are allowed to monitor and punish only half of the other
104 group members (provided the monitoring subgroups overlap). In these situations, contributions
105 are as high as when everyone monitors everyone else, but the amount of punishment received
106 by free riders is just enough to get them to contribute at high levels. That is, there is less wasted
107 punishment compared to the first technology. Lastly, when group members see and can punish
108 just one other group member, punishment is not an adequate deterrent, and contributions collapse
109 at a rate that is similar to the no-punishment control treatment.

110 There are also a few experiments that show that punishment affects free riders even when it
111 imposes no material cost on them. [Masclot et al. \(2003\)](#) show that even “cheap talk” punishment
112 is effective. Participants in the experiment assign punishment to free riders even though the
113 punishment does not reduce the target's payoff. Amazingly, free riders respond by contributing
114 more in the future. However, there is one small problem with the Masclot et al. design: it is costless
115 to mete out punishment, and therefore there are equilibria in which strictly egoistic players punish
116 along side those with preferences to punish. [Carpenter et al. \(2004a\)](#) adjust their protocol so that
117 punishment is costly to the punisher but still imposes no material harm on the target and bring
118 the resulting game to the field. Their participants are slum dwellers in southeast Asia who face
119 the sort of social dilemmas on a daily basis that the game is meant to model (e.g., clean water

120 and solid waste). Interestingly, while participants in Ho Chi Minh City and Bangkok both use
121 punishment, it only has a significant effect on contributions in Ho Chi Minh City. The authors
122 provide some evidence that these behavioral differences may be due to culture.

123 In sum, there is a lot of evidence that participants in social dilemma experiments will punish
124 free riders. It also appears that free riders respond to punishment under a variety of conditions
125 including games in which punishment imposes no material harm. Social disapproval appears to be
126 enough to motivate some free riders. However, punishment is no panacea; there are some instances
127 in which it works poorly (e.g., in small groups with weak externalities from contributions), and in
128 other cases, it is not clear that punishment always improves the social efficiency of interactions.
129 Simple communication (perhaps confounded by the social disapproval in many group members'
130 voices) appears to elicit contributions more efficiently than force, and when force is available,
131 it appears important to limit the amount of punishment that free riders receive. The following
132 experiment extends this research by asking whether punishers adhere to the law of demand.

133 3. Experimental design

134 While the following experiment is based on the voluntary contribution mechanism (Isaac et al.,
135 1984) to test whether we can explain punishment in terms of standard economic logic, we made
136 a few changes. Our changes were designed to provide us with the data to estimate the demand for
137 punishment. First, we allowed players to monitor and punish each other. Second, punishment was
138 costly to impose, and the price of punishment changed during the course of the experiment. This
139 feature allows us to estimate the price elasticity of the demand for punishment. Third, the level
140 of provision of the public good during each round determines an income for each player from
141 which players paid to punish each other. This feature allows us to estimate the income elasticity
142 of demand. Also note that because players' earnings and the price of punishment varied over
143 the course of the experiment we are able to analyze the demand for punishment using a (more
144 powerful) within-subject design.² The specifics of our experiment are as follows.

145 Define the *price of punishment*, r , as the amount a punisher must pay in experimental monetary
146 units (EMUs) to remove one EMU from the target. Our experiment was 15 periods long, and each
147 session was split into five blocks, each block lasting three periods. The price of punishment
148 varied from block to block such that $r \in \{0.25, 0.5, 1, 2, 4\}$. We ran two treatments to balance the
149 effect of changing prices. In the decreasing price treatment r equaled 4 for the first three periods,
150 meaning the punisher spent 4 EMUs to remove 1 EMU from the target, r equaled 2 in periods four
151 through six, and so on until in periods thirteen through fifteen the price was 0.25. In the increasing
152 price treatment r started at 0.25 and cycled upward to 4. Our players were randomly assigned to
153 a treatment, and we ran a total of six sessions (three for each treatment). This design resulted in
154 a total of 18 four-person groups.

155 We used the familiar *strangers* protocol (Andreoni, 1988) under which players are randomly
156 reshuffled from group to group at the beginning of each period because we wanted to control, to
157 some extent, for strategic reasons to punish. For example, players who remain in the same group
158 may perceive that their payoffs will increase if they punish free riders early on. However, if the
159 target of one's punishment is likely to be in a different group next period, participants should
160 understand that the expected benefit of punishing will be negligible. This is especially true in the
161 current experiment where each participant monitors and can punish only one other member of

² These features differ to a significant degree from the related work of Anderson and Putterman (2005).

her current group. Controlling for strategic punishment is important because doing so allows us to focus on punishment as the expression of a social preference.

The payoff function for the voluntary contribution mechanism was augmented to account for punishment.³ Imagine groups of n players, each of whom can contribute any fraction of their w EMU endowment to a public good and keep the rest. Say player i free rides at rate $0 < \sigma_i < 1$ and contributes $w(1 - \sigma_i)$ to the public good, the benefits of which are shared equally among the members of the group.

Each player's contribution was revealed to one other player in the group who could punish this person at a price of r EMUs per sanction. Let rs_{ij} be the expenditure on sanctions assigned by player i to player j , and let s_{ki} be the sanctions player i receives from player k (the instructions explicitly mentioned that $j \neq k$); then the payoff to player i is

$$\pi_i = w[\sigma_i + nm(1 - \bar{\sigma})] - rs_{ij} - s_{ki}$$

where $\bar{\sigma} \equiv (\sum \sigma_i)/n$ is the average free riding rate in the group. The variable m is the marginal per capita return on contributions to the public good (see Ledyard, 1995). In all sessions n equaled 4, m was set to 0.5, and w was 25 EMUs.⁴

Because $1/n < m < 1$ the game without punishment is a social dilemma: group incentives are at odds with individual incentives. Each contributed EMU returns only 0.5 to the contributor, which means free riding is a dominant strategy, but if $\bar{\sigma} = 1$ then everyone is free riding fully and each player's payoff is lower than it would be if everyone contributed fully. The game is finitely repeated, which implies that subgame perfection predicts free riding on every round.

Notice that adding the possibility of punishment does not change the subgame perfect prediction. Because sanctions are costly to impose and any potential benefits from getting a free rider to contribute cannot be fully internalized by the punisher, punishment is incredible and therefore cannot be a component of any subgame perfect equilibrium. Without credible punishment, free riding is still subgame perfect.

As noted above, each player monitored and was able to sanction only one other member of the group. This design feature was added to control for other possible strategic or coordination reasons that might affect players' punishing propensities. For example, if each player monitors and can punish all the other members of the group, there are at least two problematic scenarios that may arise. First, from a strategic perspective, a player may be less likely to punish a free rider because she thinks she can free ride on the punishment of others. Second, a player may be less likely to punish because she cannot explicitly coordinate her punishment efforts with the rest of her group.⁵ For example, she may feel that the free rider should be punished, but also that there is

³ The instructions (Appendix A available on the JEBO website) referred to "reductions" with no interpretation supplied.

⁴ While the current protocol is quite standard in the public goods literature in terms of group size, matching rule (partners versus strangers), and the return on contributions to the public good, the punishment and public good literature is too small yet to have any consensus on a protocol. The nexus of this literature, Fehr and Gächter used four person groups, a return of 0.4 per contribution, and examined both partners and strangers groupings. Despite this role model, more recent experiments have varied the group size between 3 and 10 members (compare Anderson and Putterman, 2003 and Carpenter et al., 2004a), or used returns on the public good between 0.3 and 0.75 (see Bowles et al., 2002; Bochet et al., 2003, and Sefton et al., 2001), and both the partners and strangers protocols are used in addition to the *complete* strangers protocol in which players are guaranteed never to be in exactly the same group twice. The most important innovation, however, has been to use the less complicated punishment mechanism described here and used in much of the more recent punishment literature. In Fehr and Gächter players paid to reduce a target's earnings by a certain percentage, implying that the "price" of punishment is not constant.

⁵ This rationale is consistent with the "unresponsive bystander" hypothesis advanced in Latane and Darley (1970).

195 an appropriate level of punishment that fits the infraction. If she does not know or cannot estimate
 196 how much others will punish, she may withhold sanctions to be sure that the punishment does
 197 not exceed the offense. If all players see only one other player and knows that the person they are
 198 monitoring is not monitoring them, we control for any strategizing and coordination problems.
 199 People should only pay to punish if they wish to express their preferences.

200 4. Overview of the data

201 We recruited 72 participants (36% were female) from the undergraduate population at Mid-
 202 dlebury College in our six experimental sessions. Participants were from a variety of majors, and
 203 most were between 18 and 22 years old. Because the experiment occurred near the end of the
 204 semester when students' opportunity cost of time is particularly high, we calibrated the game so
 205 that payments would be generous and so that we would not have any problems recruiting partic-
 206 ipants. On average, the students earned \$26.26, including a \$5 show-up fee. The typical session
 207 lasted a little less than an hour. We begin our analysis by giving the reader a broad sense of the
 208 data, and then we focus on our estimates of the demand for punishment and whether punishment
 209 is effective.

210 Reviewing previous punishment experiments (e.g., Fehr and Gächter, 2000) we see that the
 211 typical time path of contributions, averaging across treatments, starts near half the endowment
 212 and then increases at a decreasing rate. However, as seen in Carpenter (in press), punishment
 213 has less of an effect on contributions when players monitor only a small subset of their group-
 214 mates. Pooling across periods the two price change treatments exhibit nearly the same levels
 215 of contributions ($\bar{c}_{\text{increase}} = 7.55$, $\bar{c}_{\text{decrease}} = 7.59$), but such pooling does not account for the
 216 dynamics of contributions.⁶ Fig. 1 illustrates the time paths from the current experiment. In one
 217 sense, the current contributions data look similar to the same monitoring treatment of Carpenter
 218 (in press) in that punishment seems to maintain initial contribution levels, at best.⁷ However, in
 219 another sense, the current data is markedly different because contributions seem to be affected by
 220 whether the price of punishment is increasing or decreasing.

221 The treatment effect in our contributions data seems reasonable from an economic point of
 222 view. When the price of punishment starts at a relatively low level and then increases over the
 223 course of the experiment, contributions fall steadily and more dramatically than when the price
 224 is constant (as in Carpenter, in press). One explanation, which we will confirm below, is that our
 225 players based their punishment decisions on price as well as on how egregiously the target free
 226 rode. On the one hand, when the price increased over time players bought less punishment causing
 227 the threat of punishment to abate. This led to more free riding. On the other hand, when the price
 228 fell players responded by buying more punishment per offense. In this case, the effectiveness of
 229 punishment increased over the course of the experiment, and although we see an initial drop in
 230 contributions, they recover as the price of punishment continues to fall.

231 Without accounting for the time paths or the amount of free riding, the average amount of
 232 punishment purchased in the decreasing price treatment ($\bar{s}_{\text{decrease}} = 2.13$) is marginally signif-
 233 icantly greater ($t = 1.72$, $p = 0.08$) than the average amount purchased in the increasing price
 234 treatment ($\bar{s}_{\text{increase}} = 1.55$). Fig. 2 presents the time paths of the average expenditure on pun-
 235 ishment. Even though this graph does not control for other factors that might have affected our

⁶ In fact the pooled averages are not significantly different ($t = 0.09$, $p = 0.93$).

⁷ We cautiously make this comparison because, although the monitoring network was identical, the Carpenter et al. (2004a) experiment used the Fehr and Gächter payoff function.

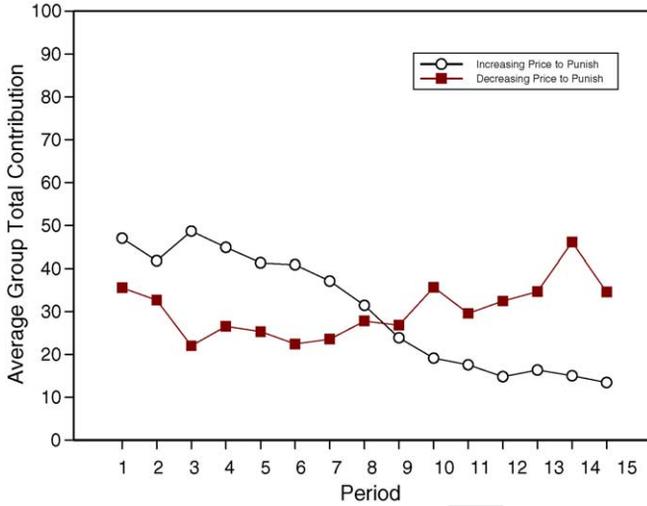


Fig. 1. The evolution of average group contributions over time (note: increasing price indicates that the price per sanction increased from 0.25 to 4 while decreasing price means the opposite).

236 players' punishment decisions (e.g., income or average level of free riding), it provides evidence
237 consistent with the hypothesis that players reacted to the price of punishment and that this affected
238 the credibility of punishment and the level of contributions. As the price increased, our players
239 spent less on punishment. In fact, by the last three rounds of the increasing price treatment when
240 it cost 4 EMUs to remove 1 EMU from the target, the players stopped punishing completely.
241 In the other treatment, as the price fell, players spent more on and bought more punishment.

242 Because the instructions explicitly mentioned the order in which the price of punishment
243 would change (see Appendix A on the JEBO website), one might worry that players anticipated

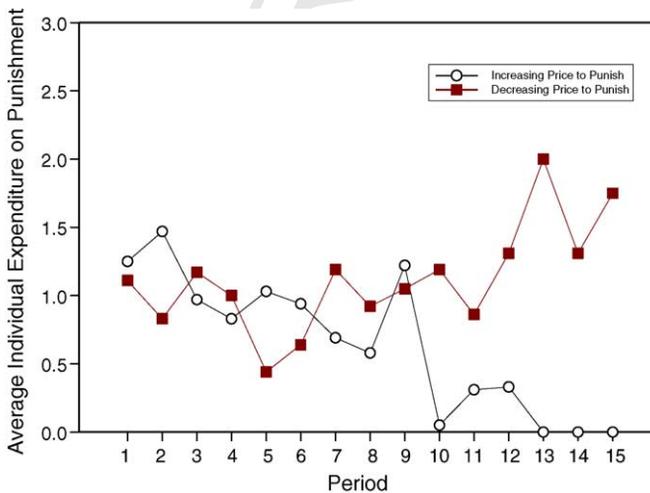


Fig. 2. The time path of punishment expenditures (note: this figure includes all the cases where players choose not to punish and does not control for how much free riding occurred).

244 and reacted in advance to the direction of the price change. For example, in the increasing price
245 treatment, it might be reasonable to think that players spent more on early punishment than
246 they would have had they not known that the price was going to increase or that players in the
247 decreasing price treatment might have delayed punishment to later rounds when they knew it
248 would be cheaper. If this is true then the slopes of the graphs in Fig. 2 are steeper than they would
249 have otherwise been. In the next section we control for these differences when estimating the
250 demand for punishment and see that they do not matter.

251 5. The demand for punishment

252 We now proceed by econometrically estimating the demand for punishment. One valuable
253 benefit of using an experiment to elicit the data for our estimation is that we control for most of
254 the problems that typically plague demand estimates. Specifically, simultaneity and identification
255 are not problems for us because price is, by design, completely exogenous. However, we do face
256 other issues. Because our experiment is 15 periods long, we generate a panel of data. To control
257 for individual heterogeneity, all our regressions include random effects. Because there are a lot
258 of observations where our players showed no preference for punishment, our dependent variable,
259 the quantity of punishment purchased, is truncated from below at zero. For this reason, we use the
260 Tobit procedure.⁸ Finally, there is one criticism of the strangers matching protocol that has been,
261 to this point, ignored in the literature. It could be the case that our point estimates will be biased
262 by the fact that observations at the individual level are not independent because contaminants
263 are generated when participants are randomly reshuffled into new groups at the beginning of
264 each round. Of course, as is true of most procedural criticisms, this is an empirical question. In
265 Appendix B (on the JEBO website) we offer a methodology for testing whether such a bias affects
266 the coefficients that we calculate below. Saving the details for the readers of the appendix, the
267 upshot is that we are confident that any strangers bias has a minimal effect on our estimates.

268 Before we present the fully controlled estimate of the demand for punishment, we present the
269 reader with a graphical presentation of the main result. Fig. 3 illustrates the uncontrolled demand
270 for punishment function based on a quadratic specification. As the reader can see, while the best
271 fit is nonlinear, the slope of the demand function is relatively shallow, indicating that the reaction
272 of participants who punish to price changes is rather small. However, there is some response,
273 and it is important to note that Fig. 3 does not represent the behavior of the entire population of
274 participants because there are a significant portion of our participants (32%) who did not punish
275 at all. We return to the issue of punishing types in Section 6.

276 We can demonstrate the robustness of Fig. 3 by considering the regression results presented in
277 Table 1. We build our econometric model in two stages. In stage one, we estimate the uncontrolled
278 price and income elasticities. In stage two we add controls for how much the target free rode (the
279 null hypothesis being one is punished more the more one free rides) and for how much the punisher
280 free rode (the null, in this case, being that people who free ride less, punish more). Our definition of
281 free riding is based on the results of Fehr and Gächter who show that people direct punishment at
282 targets who contribute less than the group average. Hence, for our purposes *free riders* contribute
283 less than the current group average, and *contributors* contribute at or above the average. In the

⁸ Technically, punishment could also be limited by a participant's earnings. However, we can report that the "budget constraint" never bound our participants' desire for punishment. In all cases the difference between a participant's period earnings and her period expenditures on punishment was positive.

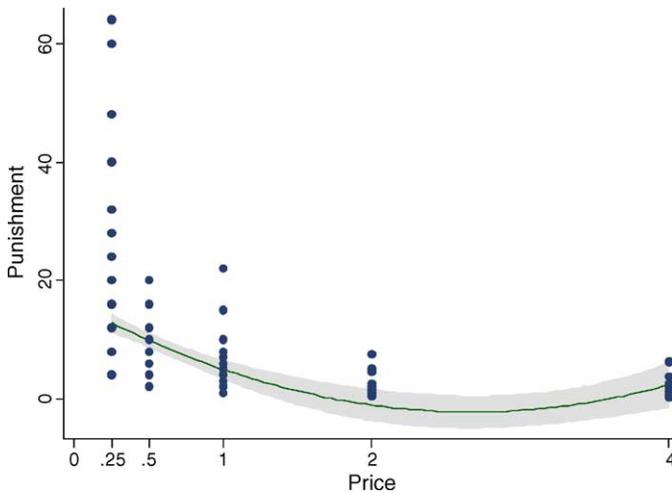


Fig. 3. The demand for punishment (note: each dot may represent several observations. The fitted values are based on a quadratic estimation of the effect of price on quantity. Shading represents the 95% confidence interval).

284 second stage we also control for the sex of the punisher, whether the punisher participated in the
 285 increasing price treatment or not, the differential effect of price changes in the increasing price
 286 treatment, and any time trend in our data. To control for the time trend we model a simple dynamic
 287 that says that the amount of punishment purchased by participant i in period t is a function of the
 288 contributions of the other group members, $C(\text{group total})_{-i,t-1}$, in period $t - 1$.⁹

289 In the upper panel of Table 1 we report the marginal effects and standard errors of our regres-
 290 sions, and in the middle panel we report elasticities calculated at the regressor means.¹⁰ Eq. (1)
 291 introduces our main result. Both the price and income elasticities are negative, which indicates
 292 that, given the average participant prefers to punish free riders, people react to economic incentives
 293 in what economists would consider reasonable fashion. In addition, demand appears to be slightly
 294 elastic with respect to price and inelastic with respect to income. Specifically, in the uncontrolled
 295 regression a 1% increase in price reduces the quantity of punishment demanded by 1.22% and a
 296 1% increase in income decreases the amount of punishment demanded by 0.27%. At first blush,
 297 punishment appears to be ordinary and inferior.

298 Eq. (2) indicates that our initial elasticity estimates are not entirely robust to the inclusion
 299 of other punishment determinants. Part of the variation in punishment previously attributed to
 300 changes in price is actually caused by changes in how egregiously the target and the punisher free
 301 ride, but the coefficient on price (and its square) remain highly significant. At the same time, there

⁹ Notice that we could have included period fixed effects as an alternative to modeling a dynamic, but this approach would be incorrect because doing so assumes there are time idiosyncrasies while what we need to control for appears to be (based on Figs. 1 and 2) a dynamic process.

¹⁰ Calculating elasticities from Tobit coefficients is not straightforward because when one calculates the marginal effect, one has to account for the probability that a change in the regressor will push one past the “kink,” and the impact of a change in the regressor on the dependent variable, given it is uncensored. However, we can use the McDonald and Moffitt (1980) decomposition to calculate elasticities. With latent variable, p_{it} we have the following marginal effect:

$$\frac{\partial E(p_{it}|x_{it})}{\partial x_{it}} = E(p_{it}^*|x_{it}, p_{it}^* > 0) \frac{\partial \Pr(p_{it}^* > 0)}{\partial x_{it}} + \Pr(p_{it}^* > 0) \frac{\partial E(p_{it}^*|x_{it}, p_{it}^* > 0)}{\partial x_{it}}.$$

Table 1

Dependent variable is the punishment inflicted on target $_{i,t}$

	(1) All types	(2) All types	(3) Free riders who punish cooperators	(4) Free riders who punish free riders	(5) Cooperators who punish free riders
Price $_{i,t}$	-3.00***(0.40)	-1.93***(0.49)	-0.10(0.63)	-3.30***(0.96)	-1.13(0.90)
Price $^2_{i,t}$	0.49***(0.09)	0.32***(0.09)	0.02(0.13)	0.52***(0.18)	0.17(0.15)
Income $_{i,t}$	-0.002***(0.0008)	-0.0007(0.001)	0.0001(0.0007)	-0.002(0.004)	-0.0004(0.002)
$C(\text{target})_{i,t} - C(\text{group average})_{i,t}$		-0.13***(0.02)	-0.0008(0.006)	-0.19***(0.05)	-0.13***(0.05)
$C(\text{punisher})_{i,t} - C(\text{group average})_{i,t}$		0.07***(0.02)	0.002(0.01)	0.04(0.04)	0.12***(0.05)
$C(\text{group total})_{-i,t-1}$		0.004(0.005)	0.001(0.005)	0.03**(0.01)	-0.007(0.006)
Female $_i$		0.10(0.40)	0.001(0.03)	-0.07(0.63)	-0.37(0.28)
Increasing price $_i$		-0.50(0.56)	0.02(0.12)	-1.27(1.39)	0.60(1.04)
Price $_{i,t} \times$ Increasing price $_i$		-0.32(0.33)	-0.11(0.48)	0.25(0.79)	-0.99(0.65)
Price elasticity	-1.22	-0.79	-0.03	-1.03	-0.34
Income elasticity	-0.27	-0.09	0.02	-0.19	-0.04
Number of censored observations	845	795	44	234	211
N	1080	1008	56	322	308
Log likelihood	-1242.65	-998.21	-51.28	-414.67	-394.99
Wald χ^2	117	237	13	88	167
p -value	< 0.01	< 0.01	0.19	< 0.01	< 0.01

Note: (1) Tobits with individual random effects. (2) Coefficients are marginal effects. (3) Elasticities calculated at the regressor means. (4) The N in (2) is lower due to differencing. (5) Columns (3)–(5) restrict attention to three of five behavioral types identified in the data. (6) * indicates 0.10, ** indicates 0.05, *** indicates 0.01.

is a dramatic change in our estimate of the income elasticity. There is a simple explanation for why the income regressor shrinks. The punisher's income is highly correlated with her deviation from the group average contribution ($\rho = 0.75$, $p < 0.01$). Without controlling for the punisher's level of free riding, the income regressor picks up the variation due to both how much the punisher free rides and how much income is generated by her group.

Controlling for the other determinants of punishment, we continue to find that punishment is inferior based on the point estimate; however this estimate is actually not significantly different from zero, indicating that experimental income changes have little effect on the decision to punish. With respect to price, we find that punishment remains ordinary, but it is now less responsive to price. A one percent increase in the price of punishment leads to a 0.79% reduction in the quantity demanded. We conclude that the demand for punishment is ordinary, inferior, but inelastic.¹¹

Eq. (2) also reveals other interesting facts about punishment, only some of which have been documented elsewhere. As is now common, we find that punishment decreases as the target increases her contributions ($p < 0.01$) and punishment is meted out primarily by people who themselves contribute a lot ($p < 0.01$). As for results one does not typically see in this literature, we find that the lag of the contributions by one's group mates has little effect on punishment decisions, females purchase more punishment (but not significantly more), and the increasing price treatment does not have a direct effect on the quantity of punishment purchased nor does it have a significant differential effect on the impact of price.

6. Types of punishers

Experimenters are becoming as interested in the heterogeneity of behavior as they used to be in average behavior (e.g., Fischbacher et al., 2001). Given this interest, we notice that there are six basic "punisher types" that can be found in our game. On one dimension players can contribute or not, and on another dimension, players can either not punish, punish free riders, or punish contributors. To implement this classification system, we sort players by their average contribution over the 15 periods (identifying those who contribute less than the average as free riders), and then for each individual, we regress the amount of punishment a player bought on the contribution of their target to get an estimate of each player's punishment propensity.¹² We find support for five of the six possible types; there are no players who contribute themselves and punish other contributors. The distribution of the five supported types is as follows: 5.56% are free riders who punish cooperators, 16.67% are free riders who never punish, 31.94% are free riders who punish other free riders, 15.28% are contributors who do not punish, and the remaining 30.56% are contributors who punish free riders. We refer to the first group as *principled free riders*, the third group as *hypocritical free riders*, and the last group as *principled cooperators*.

¹¹ Averaging across periods, treatments, and individuals those participants who punished spent 14% of their per period income on sanctions. Given this is a significant fraction of their earnings, a change in the price of punishment has a dramatic effect on their *real* budget constraints in the experiment. For this reason, we also decomposed our data into an income effect that is not picked up in our income elasticity, and a pure substitution effect. The relative size of the substitution effect accounts for more than 99% of the observed change. In other words, the income effect is negligible and the substitution effect is large.

¹² There were actually plenty of degrees of freedom to conduct these regressions because each individual generates 15 observations.

338 Returning to Table 1, we can assess whether individual types react differently to changes in
339 price and income (as well as to the other determinants of punishment). In an unreported regression
340 we stacked the data and created interaction terms to test whether any coefficient differences that
341 we see in Table 1 are significant. At a minimum, the test of whether all the coefficients on
342 the interactions were jointly different from zero is highly significant ($\chi^2 = 111.82$, $p < 0.01$),
343 indicating some differences are important.

344 Notice that in the analysis of punishment, only three of the five types are important because
345 two types never punish. First, in Eq. (3) we see that free riders who punish cooperators do not
346 condition their punishment on any of the determinants of punishment. The fact that none of
347 the regressors predict the behavior of this type of punisher may simply be due to the fact that
348 we have few observations in this cell of the design. Alternatively, these principled free riders
349 may feel as if they have “figured out” the game and are using punishment to indignantly signal
350 to high contributors that they are being foolish. Another large portion of the players are
351 hypocritical. In Eq. (4) we see that these hypocrites (those who free ride and punish other
352 free riders) are more sensitive to price changes than any of the other types. The stacked regression
353 suggests that our hypocrites are more sensitive to price changes than vengeful free
354 riders, but the effect is only at the margins of significance ($p = 0.17$ for the first order comparison
355 and $p = 0.16$ for the second order comparison). However, hypocrites are significantly
356 more sensitive to price ($p < 0.10$ for both coefficients) than cooperators who punish free riders
357 (Eq. (5)). It makes sense that although there may be some trace of morality to the punishment
358 choices of the hypocrites, they are much more likely to shed their morals when the
359 cost of punishing increases. At the same time, principled free riders and principled cooperators
360 blindly follow established cultural norms (i.e., their responses are not significantly conditioned
361 on the cost of punishment), albeit the norm followed is pro-social in only one of the two
362 cases.

363 The other significant coefficients in Eqs. (3)–(5) provide more evidence supporting this normative
364 story. Like principled cooperators, hypocrites condition their punishment on how much
365 the target free rides while principled free riders do not. In addition, in Eq. (5) we see that those
366 cooperators who are further above the group average punish more, but more interestingly, in Eq.
367 (4) we see that only hypocrites are sensitive to the dynamic in contributions. The coefficient on
368 the lag of the contributions of the others is positive (and significantly greater than the cooperator
369 coefficient, $p = 0.03$), indicating that the more the others contributed in the past, the more hypocrites
370 punish in the future. In sum, we find three types of punishers in our data: hypocritical free
371 riders who are very sensitive to the parameters of the game and the reactions of the other players,
372 principled cooperators who mostly condition their punishment on how badly a target free rides
373 and care little for the payoff implications of punishing, and principled free riders who punish
374 (almost indiscriminately) other group members, maybe because they want to express indignation
375 for contributors.

376 7. Does punishment deter free riding?

377 Although it seems like an aside, to understand better the elasticities we calculated in Section
378 5, it is important to ask whether punishment is effective at attenuating free riding. This question
379 is especially important if it turns out that punishment is not effective. In the pooled data, we find
380 that the demand for punishment is ordinary, but inelastic, which suggests that the dominant force
381 in our aggregate data is the behavior of the principled cooperators. We suspect these players of
382 subscribing to a norm that compels them to sanction other players who are not acting in ways

383 that contribute to group welfare. In other words (i.e., those of Elster, 1989) these players are not
384 punishing for instrumental reasons. Instead principled cooperators punish for normative reasons,
385 and as a result, they are less sensitive to the payoff implications of punishment.¹³ Specifically,
386 they do not punish to increase group welfare per se, they punish without much regard to the
387 cost of doing so, and they punish even if it does not cause free riders to contribute more in the
388 future.¹⁴

389 To examine whether punishment is effective in the current experiment, in Table 2 we look at
390 the regression of individual contributions by player i in period t , $C(\text{individual})_{i,t}$, on the amount
391 of punishment that this person received in period $t - 1$. We include other regressors such as
392 how much the other group members contributed last period to separate the effect of inertia from
393 the effect of punishment, the player's sex, and controls for the direction of price changes and
394 any differential effect of the price treatment on received punishment. We pool the data across
395 punisher types in Eq. (1) and find that punishment has no effect on future contributions.¹⁵ In
396 fact, the only regressor with any significant explanatory power is $C(\text{group total})_{-i,t-1}$, which
397 implies that increases in contributions in this experiment can be attributed to inertia, conform-
398 ity (Carpenter, 2004), or conditional cooperation (Fehr and Fischbacher, 2003). Given the
399 tiny coefficient on the lagged punishment, it must have been obvious to the participants that
400 punishment did not deter free riding, yet as we see in Fig. 2, participants continued to pur-
401 chase sanctions. As mentioned above, this can not be reconciled with the strategic use of
402 punishment, but it is consistent with the normative motivations that we ascribe to principled
403 cooperators.

404 In Eqs. (2) through (6) of Table 2, we dig a little deeper and examine the effect of being
405 punished on our five behavioral types individually. Inertia seems to be the only common de-
406 terminant of cooperation, although different types respond differently to the contributions of
407 others. It is interesting, for instance, that the inertial coefficients on the two cooperative types
408 are three times as large as the coefficients on the three free riding types. These differences, sig-
409 nificant in the stacked regression at the 2% level or better, suggest that cooperators are three
410 times as motivated by conditional cooperation as free riders are; however, none of the pun-
411 isher types are significantly motivated to contribute when, instead of inflicting it, they receive
412 punishment.

413 There are two other interesting facts in the type-level regressions that are worth mentioning.
414 Women who are vengeful free riders are also significantly more stubborn in their free riding. While
415 not significant, the coefficient on the lag punishment regressor is negative, suggesting punishment
416 makes these women free ride even more. This effect occurs in addition to the sex indicator being
417 highly significantly negative with a very large marginal effect. This is the only substantial sex
418 difference we find in our data.

419 Lastly, we see that principled cooperators (Eq. (6)) have significantly lower contributions in the
420 increasing price treatment. A possible explanation for this coefficient is that principled cooperators
421 “give up” at some point in the increasing price treatment. Early on they use punishment to repair
422 the cracks in the dam that holds back the swell of free riding, but as the price of punishment

¹³ Also see Carpenter and Matthews (2005) for a one-shot experiment that reinforces the normative interpretation of punishment.

¹⁴ Actually, this model of behavior is not just speculation. For other evidence supporting this hypothesis, see Carpenter et al. (2004b), and for an experiment designed explicitly to test this idea see Carpenter and Matthews (2002).

¹⁵ Finding that punishment is ineffective in a network in which each group member monitors and can punish only one other group member replicates the results of a similar treatment in Carpenter et al. (2004a).

Table 2
Dependent variable is $C(\text{individual})_{i,t}$

	(1) All types	(2) Free riders who punish cooperators	(3) Free riders who do not punish	(4) Free riders who punish free riders	(5) Cooperators who do not punish	(6) Cooperators who punish free riders
$C(\text{group total})_{i,t-1}$	0.17***(0.01)	0.09*(0.05)	0.05***(0.02)	0.10***(0.02)	0.19***(0.04)	0.27***(0.02)
Punishment received $_{i,t-1}$	-0.06(0.10)	-0.41(1.01)	-0.13(0.12)	0.14(0.17)	0.56(0.65)	0.05(0.17)
Female $_i$	-0.28(1.40)	-3.90**(1.58)	-0.95(0.76)	0.10(0.94)	-1.00(2.98)	-1.49(1.08)
Increasing price $_i$	-0.59(1.12)	1.16(1.72)	-2.91***(0.48)	0.96(0.97)	3.63(3.12)	-3.52***(1.13)
Punishment received $_{i,t-1} \times$ Increasing price $_i$	-0.15(0.19)	0.20(1.14)	0.19(0.18)	-0.05(0.27)	-0.46(0.91)	-0.36(0.57)
Number of censored observations	376	26	102	124	44	80
N	1008	56	168	322	154	308
Log likelihood	-2553.54	-129.78	-265.37	-786.82	-445.06	-445.06
Wald χ^2	247	12	200	41	27	162
p -value	< 0.01	0.04	< 0.01	< 0.01	< 0.01	< 0.01

Note: (1) Tobits with individual random effects. (2) Coefficients are marginal effects. (3) Columns (2)–(5) restrict attention to one of five behavioral types identified in the data. (4) * indicates 0.10, ** indicates 0.05, *** indicates 0.01.

423 increases, it becomes increasingly costly to keep the dam together, and at some point they simply
424 get out of the way and stop contributing themselves.

425 8. Concluding remarks

426 At the beginning of this paper we pointed out that while laboratory experiments in economics
427 have provided more puzzles than answers, we should not be too quick to conclude that the standard
428 methodology of economics is inherently flawed. The results of the current experiment give us
429 hope that after documenting and understanding anomalies such as social preferences, economic
430 tools will remain informative. With this in mind, our analysis has demonstrated three things.
431 One, we have replicated and extended the experiments suggesting that the average economic
432 decision-maker will, at some personal cost, punish free riders who reduce the social efficiency
433 of group interactions. Adding the current evidence to that of a number of other experiments
434 illustrates that positing a preference for punishing free riders appears to be a reasonable addi-
435 tion to standard selfish preferences. Two, given we accept that people prefer to punish free
436 riders, we have shown that the most basic economic analysis, the estimation of demand, illus-
437 trates that people react to price and income changes when they consider punishing free riders
438 just as they react to changes in these variables when they consume more standard commodi-
439 ties. Specifically, the demand for punishment slopes downward and is relatively inelastic with
440 respect to price and income. If punishment preferences are linked to normative behavior, then
441 it makes sense that punishing behavior is relatively inelastic with respect to price and income
442 because people punish primarily for social rather than economic reasons. Third, despite the rela-
443 tive inelasticity of the demand for punishment, we have shown that punishers are sensitive to the
444 price of punishment but not sensitive to income changes that should allow one to punish more
445 severely.

446 These results also dovetail nicely with some of the other experimental studies of mutual moni-
447 toring. While there have been a number of recent studies that have looked at the effects of different
448 punishment mechanisms (see Section 2 and Decker et al., 2003), there has been little compar-
449 ative static analysis. For example, we know next to nothing about the robustness of any given
450 punishment scheme. For example, does the original Fehr and Gächter mechanism continue to
451 elicit contributions if each punishment point removes 5% of the target's income instead of 10%?
452 Likewise, while Sefton et al. find that punishment is better at controlling free riding than rewards,
453 how much does this result depend on the relative magnitudes of punishments and rewards? Un-
454 til there has been a systematic study of the determinants of punishing behavior similar to what
455 Ledyard (1995) has done for public goods experiments, the punishment literature will remain a
456 series of unconnected islands. The current experiment, however, extends previous work in which
457 we fixed the punishment mechanism and varied the amount of information that group members
458 have about each other (Carpenter, in press), the return on the public good, and the size of the
459 groups (Carpenter, in press and Bowles et al., 2001).

460 The fact that many punishers in our experiment react little to the monetary consequences of
461 their actions overlaps with similar situations in the real world including the interesting example
462 of mutual monitoring among lobster fishermen along the Maine coast documented in Acheson
463 (1988). In this example, fishermen monitor and punish others who extract too much from the
464 local fishery, and when they do, they risk heavy monetary fines and imprisonment because their
465 vigilante methods are often extreme and illegal (ranging from cutting trap lines to blowing up
466 boats). In this sense, our finding that punishment is inelastic with respect to price is economically
significant even if the exact estimate of the elasticity has little external validity.

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472 Appendix A. Participant instructions

473 You have been asked to participate in an experiment. For participating today and being on time
474 you have been paid \$5. You may earn an additional amount of money depending on your decisions
475 in the experiment. This money will be paid to you, in cash, at the end of the experiment. When
476 you click the BEGIN button you will be asked for some personal information. After everyone
477 enters this information we will start the instructions for the experiment.

478 During the experiment we will speak in terms of Experimental Monetary Units (EMUs) instead
479 of Dollars. Your payoffs will be calculated in terms of EMUs and then translated at the end of the
480 experiment into dollars at the following rate: 25 EMUs = 1 Dollar.

481 In addition to the \$5.00 show-up fee, each participant receives a lump sum payment of 10 EMUs
482 at the beginning of the experiment.

483 The experiment is divided into 15 different periods. In each period participants are divided
484 into groups of 4. The composition of the groups will change randomly at the beginning of each
485 period. This means that in each period your group will consist of different participants.

486 Each period of the experiment has two stages.

487 A.1. Stage one

488 At the beginning of every period participants receive a 25 EMU endowment. In stage one
489 participants decide how much of their 25 EMUs to contribute to a group project and how much to
490 keep for themselves. Participants' payoffs are determined by the total contribution of their specific
491 group and how much they individually keep.

492 To record their decisions, participants will type EMU amounts in two text-input boxes, one
493 for the group project labeled GROUP ALLOCATION and one for themselves labeled PRIVATE
494 ALLOCATION. These boxes will be yellow. Once a participant makes a decision, he or she will
495 record his or her decision by clicking on the green SUBMIT button.

496 After all the participants have made their decisions, you will each be informed of your gross
497 earnings for the period.

498 Participant gross earnings will consist of two parts:

- 499 (1) Earnings from the private allocation. Individuals are the only beneficiary of EMUs they keep.
500 Specifically, each EMU a participant keeps increases that person's earnings by one.
501 (2) Earnings from the group project. Each member of a group gets the same payoff from the group
502 project regardless of how much he or she contributed. The payoff from the group project is
503 calculated by multiplying 0.5 times the total EMUs contributed by the members of the group.

504 Participant gross earnings can be summarized as follows:

505
$$1 \times (\text{EMUs you keep}) + 0.5 \times (\text{Total EMUs contributed by your group})$$

506 Let's discuss three examples.

507 **Example 1.** Say each member of a group contributes 15 of the 25 EMUs. In this case, the group
 508 total contribution to the project is $4 \times 15 = 60$ EMUs. Each group member earns $0.5 \times 60 =$
 509 30 EMUs from the project. The gross earnings of each member will then be the number of EMUs
 510 kept, $25 - 15 = 10$, plus the earnings from the group project, 30 EMUs, for each member. In total,
 511 each member would earn $10 + 30 = 40$ EMUs.

512 **Example 2.** Now say everyone in the group contributes 5 EMUs. Here the group total contribution
 513 will be 20 and each member will earn $0.5 \times 20 = 10$ EMUs from the group project. This means
 514 that the total earnings of each member of the group will be 20 (the number of EMUs kept) plus
 515 10 (earnings from the group project) which equals 30 EMUs.

516 **Example 3.** Finally, say three group members contribute all their EMUs and one contributes
 517 none. In this case, the group total contribution to the project is $3 \times 25 = 75$ EMUs. Each group
 518 member earns $0.5 \times 75 = 37.5$ EMUs from the project. The three members who contributed
 519 everything will earn $0 + 37.5 = 37.5$ EMUs and the one member who contributed nothing will
 520 earn $25 + 37.5 = 62.5$ EMUs.

521 A.2. Stage two

522 In stage two participants will be shown the allocation decision made by one other randomly
 523 selected member of their group. Everyone's choice will be seen by exactly one other group member
 524 and the person you see is different from the person seeing you. In addition to seeing another group
 525 member's choice, at this stage participants can reduce the earnings of the group member they see,
 526 if they want to.

527 Participants will be shown how much one member of their group kept and how much this
 528 person allocated to the group project. Participants will also see their own allocation decision and
 529 this decision will be labeled 'YOU'.

530 At this point participants will decide how much (if at all) they wish to reduce the earnings of
 531 the other group member they are seeing. Participants reduce someone's earnings by typing the
 532 number of EMUs they wish to spend to reduce that person's earnings into the input-text box that
 533 appears below the other group member's allocation decision.

534 Participants can spend as much of their accumulated earnings as they want to reduce the
 535 earnings of the other group member. For each EMU spent by a participant the earnings of the
 536 other group member will be reduced by R EMUs. The value of R will change during the experiment.

537 [Price decrease] The experiment is divided into 5 blocks of 3 periods and the value of R will
 538 change every 3 periods according to the following sequence $\{0.25, 0.5, 1, 2, 4\}$. For example,
 539 during the first 3 periods of the experiment R will be 0.25 so spending 1 EMU will reduce the
 540 other group member's earnings by 0.25 EMUs. During the third block of periods R will equal
 541 1 and spending 1 EMU will reduce the other group member's earnings by 1 EMU. During the
 542 final block R will equal 4 and spending 1 EMU will reduce the other group member's earnings by
 543 4 EMUs.

544 [Price increase] The experiment is divided into 5 blocks of 3 periods and the value of R will
 545 change every 3 periods according to the following sequence $\{4, 2, 1, 0.5, 0.25\}$. For example,
 546 during the first 3 periods of the experiment R will be 4 so spending 1 EMU will reduce the
 547 other group member's earnings by 4 EMUs. During the third block of periods R will equal 1
 548 and spending 1 EMU will reduce the other group member's earnings by 1 EMU. During the final

549 block R will equal 0.25 and spending 1 EMU will reduce the other group member's earnings by
550 0.25 EMUs.

551 Consider this example: suppose someone spends 2 EMUs to reduce the earnings of the other
552 group member when R is 0.5. This expenditure reduces the other group member's earnings by
553 1 EMU ($2 \times 0.5 = 1$). When participants have finished stage two they will click the blue DONE
554 button.

555 Participant Net Earnings in each period will be calculated as follows:

556 (Gross earnings from stage one) – (R times the number of EMUs spent on reductions directed
557 towards the participant) – (the participant's expenditure on reductions directed at someone
558 else).

559 If you have any questions please raise your hand. Otherwise, click the red FINISHED button
560 when you are done reading.

561 Appendix B. Assessing any strangers bias

562 In the *strangers* matching protocol all the participants in an experimental session are randomly
563 re-matched into groups at the beginning of every decision-making period. This protocol is partic-
564 ularly valuable as a tool to gather more observations with the same number of participants as well
565 as to examine the effects of learning while controlling, as much as is practical, for repeated game
566 effects. Even though most experiments have end points that are common knowledge, the typical
567 participant finds the difference between finitely repeated play and infinitely repeated play much
568 less compelling than theorists do. As a practical matter this means that participants who interact in
569 the same group (a.k.a., the *partners* protocol), despite knowing the exact number of interactions,
570 fail to undertake the necessary backward induction required to see through pseudo-folk theorem
571 like reasoning. As a result, researchers rely on the strangers protocol to come as close to a series
572 of one-shot encounters as possible. However, as many researchers (and referees) have pointed
573 out, when it comes to econometric analysis, the strangers protocol may cause violations of the
574 assumption that observations at the individual level are independent. Taken to the extreme, such
575 a critique implies that one can only conduct analyses at the session level because this is the level
576 at which independence is guaranteed.

577 At first blush, this argument seems both correct and paralyzing for experimental research
578 because either one needs to conduct a huge number of sessions or create protocols that can
579 run with as few participants as possible. It is obvious that increasing the cost of experimental
580 research by running many more sessions is a problem, but one should not disregard the incentive
581 to run experiments on smaller groups in smaller sessions because there are many experiments
582 for which group size and anonymity matter. However, as is true about most questions concerning
583 experimental methodology, this is really an empirical question. While the logic is sound, if such
584 dependence among individual decisions leads to no or only minor bias in the important point
585 estimates, then it seems imprudent to neglect analyses at the individual level. To this point, most
586 researchers simply ignore this stranger's bias and hope that their referees will not call them on
587 it. However, in what follows, we offer two methods for assessing the magnitude of this potential
588 bias.

589 In Table B.1 we report the results of two tests of the stranger's bias in our demand for punish-
590 ment data. In columns (1) through (3) we compare three different estimates using our punishment
591 data. Notice that these regressions do not contain all the same variables as our preferred model in

Table B.1

Dependent variable is the punishment inflicted on target $_{i,t}$

	(1) Tobit	(2) Tobit	(3) Tobit	(4) GLLAMM	(5) GLS
Price $_{i,t}$				−4.44*** (0.54)	−4.46*** (0.54)
Price $_{i,t}^2$				0.78*** (0.11)	0.78*** (0.11)
Income $_{i,t}$	0.36(0.44)	0.39*** (0.15)	0.61*** (0.13)	0.001(0.003)	0.001(0.002)
$C(\text{target})_{i,t} - C(\text{group average})_{i,t}$	−1.32*** (0.49)	−1.01*** (0.12)	−1.02*** (0.12)	−0.23*** (0.03)	−0.22*** (0.03)
$C(\text{punisher})_{i,t} - C(\text{group average})_{i,t}$	0.15(0.60)	0.97*** (0.21)	1.26*** (0.18)	0.13*** (0.03)	0.13*** (0.03)
$C(\text{group total})_{i,t} - C(\text{group total})_{i,t-1}$		0.10*** (0.04)		0.02* (0.01)	0.02* (0.01)
Female $_i$	2.18(4.51)	−0.09(2.48)	−0.03(2.55)	−0.31(0.50)	−0.31(0.48)
Increasing price $_i$				−1.09(0.90)	−1.13(0.89)
Price $_{i,t} \times$ Increasing price $_i$				0.23(0.51)	0.25(0.51)
Number of censored observations	50	795	795		
N	72	1008	1008	1008	1008
Log likelihood	−109.71	−1070.77	−1074.25	−2999.78	
Wald χ^2	9	149	143		292
p -value	0.06	< 0.01	< 0.01		< 0.01

Note: (1) Eqs. (1)–(3) are Tobits censored at 0. (2) The GLLAMM procedure stands for generalized linear latent and mixed models and incorporates random effects at both the individual and session levels. (3) The generalized least squares (GLS) model incorporates individual random effects. (4) * indicates 0.10, ** indicates 0.05, *** indicates 0.01.

592 **Table 1.** We do not include any of the price variables because they do not vary over their entire
 593 range in period one. For example, the first order price regressor takes only two values, 0.25 or 4. In
 594 column (1) we estimate the relationship using just the data from period one that implies there can be
 595 no stranger's bias affecting the coefficients. In columns (2) and (3) we use the data from periods two
 596 through fifteen. If the random reshuffling of partners causes problems in our data, we should be able
 597 to identify whether or not the effect is large enough to bias our point estimates by comparing the co-
 598 efficients based on the period one data to the coefficients based on the remaining data. This is a sim-
 599 ple Hausman test. What complicates the matter, however, is that learning may occur over the course
 600 of fifteen periods. This is why we compare column (1) to columns (2) and (3). Columns (2) and (3)
 601 differ in that column (2) controls (in an admittedly simplistic, but agnostic way) for learning by
 602 adding the lagged contributions of the other group members. As one can see, the coefficients look
 603 similar, but not exactly the same so the Hausman results are important. Comparing columns (1)
 604 and (3), which does not include a control for learning, the p -value on the Hausman test is 0.21 and
 605 we can not reject the hypothesis that the coefficients are the same (i.e., there is minimal stranger's
 606 bias). This result does not change much when we add the learning control. In this case we find
 607 $p = 0.27$. Based on this method, we conclude that the stranger's bias is not important in our data.

608 A second way to identify an effect of random re-matching (in which we use all our regressors)
 609 is to compare the standard method of analysis used in this literature, generalized least squares
 610 with individual random effects, which ignores any potential stranger's bias, to the generalized
 611 linear latent and mixed models (GLLAMM) approach. In the GLLAMM procedure we include
 612 individual random effects and session level random effects to capture the stranger's bias. As one
 613 can see, the difference in the coefficients between the GLLAMM model (column (4)) and the GLS
 614 model (column (5)) are almost imperceptible, which provides further evidence that the stranger's
 615 bias does not affect the coefficients of our estimates of the demand for punishment. Unfortunately,
 616 so far a GLLAMM model has not been developed for the Tobit regressor, but given the results of
 617 our two methods for assessing the magnitude of the stranger's bias in our data, we have confidence
 618 in the random effect Tobit regressions reported above.

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